The Value Of Item Response Theory In Invariance Testing

Nico Martins, University of South Africa, South Africa

ABSTRACT

The goal of the current study was to assess the Employee Engagement Instrument (EEI) from an item response theory (IRT) perspective, with a specific focus on measurement invariance for annual turnover. The sample comprised 4 099 respondents from all business sectors in South Africa. This article describes the logic and procedures used to test for factorial invariance across groups in the context of construct validation. The procedures included testing for configural and metric invariance in the framework of multiple-group confirmatory factor analysis (CFA).

The results confirmed the factor analytic structure of the model fit for some of the individual scales of the EEI. The measurement invariance of the EEI as a function of annual turnover was confirmed. However, the results indicated that the EEI needs to be refined for future research.

Keywords: Invariance; Employee Engagement; Measurement Equivalence; Configural Invariance

INTRODUCTION

he validation of assessment instruments is a vital prerequisite in most fields such as psychology, education, training, human resources and other related behavioural fields. In most cases, assessment instruments are tested in a single organisation or in different organisations in one country (Abrams, Lee, Brown, & Carr, 2015; Martins & Da Veiga, 2015; Martins, 2014; Moerdyk, 2009) for validity and reliability for the total population. The validation of assessment instruments is generally not applied among homogeneous groups (i.e. one group with the same gender, language, qualifications or national culture) or the same context (see Egri, 2013; Fahr, Cannella, & Lee, 2006; Johns, 2006; Rousseau & Fried, 2001; Tsui, 2006; Whetten, 2009). When comparing groups, researchers often assume that the assessment instrument measures the same psychological construct in all groups or operates exactly in the same way for each group under investigation (Byrne, 2004; Dimitrov, 2010; Martins, 2014; Milfont & Fischer, 2010; Nel & Boshof, 2014; Rock, Werts, & Flaugher, 1978; Tay, Meade, & Cao, 2015). Milfont and Fischer (2010: 112) summarise this research dilemma as follows: "Despite its appeal, this assumption is often not justified and needs to be tested." According to Meiring, Van de Vijver, Rothman, and Barrick (2005) and Moerdyk (2009), South African researchers have to consider the diverse South African population in measurement. This observation is supported by many studies reporting on measurement equivalence (ME) testing across different groups, which concluded either partial invariance or non-invariance (e.g. see Dawson, Sotelo, Roesch & Klonoff, 2014; González & Jenkins, 2014; Horn & McArdle, 1992; Spurk, Abele, & Volmer, 2015). According to Strasheim (2011: 38) and Steenkamp and Baumgartner (1998:78), using instruments that are not invariant across groups may lead to invalid findings, which may limit the usefulness of the study or measurement.

Currently, two methods are commonly used for assessing measurement equivalence/invariance. One method is based on confirmatory factor analysis (CFA) and the other on item response theory (IRT) (Meade & Lautenschlarer, 2004: 361; Raju, Laffitte, & Byrne, 2002: 517; Reise, Widaman, & Pugh, 1993: 552). According to some authors (Raju et al., 2002; Tay et al., 2015:4), IRT can provide different and potentially more useful information for the establishment of measurement invariance. Several articles are available in the literature discussing the various methods (Abrams et al., 2015; Embretson & Reise, 2000; Hays, Morales, & Reise, 2000, Dimitrov, 2010; Makransky & Bilenberg, 2014; Steenkamp & Baumgartner, 1998; Vandenberg & Lance, 2000); or the similarities

and differences between these methods, as well as when it is best to use each method and the many different ways in which each method can be applied (e.g. see Meade & Lautenschlarer, 2004; Raju et al., 2002; Reise et al., 1993). Tay et al. (2015) recommend that researchers use both CFA and IRT whenever possible to ensure the best results.

A methodology for testing assessment instruments across groups is the testing of factorial invariance (Reio & Shuck, 2015). According to Dimitrov (2010), testing for factorial invariance across groups is often loosely applied (or not at all) in studies dealing with the validation of assessment instruments in counselling and education. Researchers who test for invariance such as Samuel, South, and Griffin (2015) explain the need to examine for measurement equivalence, or invariance, which would indicate that the scores on the measure have the same meanings for men and women (i.e. different groups) in validating the Five Factor Personality Scale. La Greca, Ingles, Lai, and Marco (2015) also applied factorial invariance to examine the factorial invariance and latent mean differences of the Social Anxiety Scale for Adolescents across gender and age groups in Hispanic American adolescents. Their findings support Moerdyk's (2009) observation that it is especially important in a multicultural country such as South Africa, with its numerous language and ethnic groups, to take differences into account in order to conduct fair assessments.

BACKGROUND

Employee engagement is a much-debated topic among people across business, academia and government (Imandin, Bisschoff, & Botha, 2014; Martins & Ledimo, 2016; Robbinson & Gifford, 2014). A study of nearly 8 000 business units in 36 companies by Harter, Schmidt, and Hayes (2002) found that those companies whose employees had high average levels of engagement also had higher levels of customer satisfaction, were more productive, brought in higher profits and had lower levels of staff turnover and accidents than those at other companies. These positive perceptions of employee engagement were also reflected by Robbinson and Gifford (2014) in their study. Some of the contributors argued (p. 1) "that employee engagement brings big benefits to organisations. It is an excellent vehicle for driving culture change, and is good for employees, too. However, employee engagement is hard work and requires more than a one-off 'initiative'; sustained effort is needed over a long period of time, and there are no 'quick fixes' ". Others stated (p. 3) that "as a believer in the whole debate about the importance of engagement and business performance, ... [to] me, it has always seemed obvious that the way a workforce feels about their place of work will materially affect the performance of that organisation. This is particularly relevant in service organisations, where customers are at the receiving end of good or bad attitudes. But, it is also relevant where businesses make things, or exist for other purposes." In light of the introduction, it makes business sense to ensure that the assessment tool applied to measure employee engagement is not only valid and reliable for a population, but also portrays factorial invariance across the groups included in an assessment to ensure that researchers have confidence that they are measuring what they are supposed (or claim) to measure.

This leads to the aim of this research, namely to assess the Employee Engagement Instrument (EEI) with a specific focus on measurement invariance for annual turnover. Annual turnover is generally linked to organisational size and the classification of small, medium and large organisations, reflecting different contexts. The South African market structure can best be described as an oligopoly because a few large firms dominate the market. Despite the oligopolistic nature of the South African market, small, medium and micro (SMMEs) businesses form avital part of the economy. Moreover, the interests of SMMEs are facilitated by the provisions of the National Small Business Act 102 of 1996 (as amended). This Act defines SMMEs in terms of a range of measures, including annual turnover. According to the classification of this Act, the turnover of micro enterprises is generally less than R0.2 million per annum, while the turnover of medium enterprises does not exceed R64 million. Hence, in terms of these definitions, an enterprise with a turnover of more of than R64 million can be considered large. As such, the turnover guidelines of this Act can also be used for purposes of this study. In light of the profile of the respondents, the turnover category of more than R6 million but less than R32 million is considered to represent small enterprises, while the turnover category of more than R32 million but less than R64 million is deemed to represent medium enterprises. Moreover, organisations of different sizes may have different employee value propositions and thus may attract employees with different attributes, skills sets and attitudes. Smaller organisations may also be more susceptible to economic cycles - for example, the global financial crisis, to mention a few possible differences. Employees working in different organisational contexts may thus experience engagement differently. Hence turnover is a useful

measure to determine if the construct of engagement is invariant across different business contexts, as reflected by annual turnover.

As such, this study might provide deeper insight into the measurement of engagement of employees across different size organisations that might be significant for theoretical advances in employee engagement research and for practitioners in particular. This is one of the first studies to report on the testing of ME for engagement, in particular in a multicultural society.

RESEARCH OBJECTIVES

The main objectives of this study were as follows: (1) to confirm the factor analytic structure of the measure; (2) to determine the model fit for the individual scales of the EEI; and (3) to determine the measurement invariance of the EEI as a function of annual turnover.

Research Approach

The researchers initially applied exploratory factor analysis to evaluate the factorial structure of the instrument and Cronbach's alpha to determine/ establish the internal reliability of the scale and subscales. The internal reliability and construct validity were again confirmed in a second study by means of confirmatory factor analysis (CFA) (Nienaber & Martins, 2015). In order to further unpack the questionnaire's factorial analytic structure, it was decided to conduct an IRT analysis. IRT can be described as a collection of models designed to statistically analyse items in the scale for the purposes of test development, validation and scoring (source). It is considered a technique that offers several advantages over the traditional classical test theory (CTT) approach (Ho Yu, 2010), including the following:

- In CTT, the unit of analysis is typically an aggregate (e.g. a summed or mean test score). IRT, as the name suggests, focuses on an item level and assesses how said item is related to an estimated latent trait (Fan, 1998).
- Owing to the fact that IRT calculations are said to be sample or test independent, item parameters can be compared across samples. Conversely, CTT calculates an individual's score in the context of a specific questionnaire and item properties only hold true for that sample.
- IRT also calculates item-independent respondent measures. This implies that a group of test takers can complete different sets of items, yet be measured on the same scale of performance.
- As an item level analysis technique, IRT can provide a measurement of individual error, as opposed to CTT's total standard error of measurement.

In addition to the advantages outlined above, there is also an important consideration that relates specifically to the Likert scale categorical response format of the EEI. In CTT, Likert scale data is often assumed and treated as an interval or ratio level of measurement, and unit increases in the rating scale are assumed to be equal (Bond & Fox, 2007). Many researchers find this assumption worrisome, given the fact that an equal distance between numbers is not a true reflection of Likert data (Linacre, 2005). If we hold Likert scale items to the ordinal level of measurement, then this would render parametric statistics untenable.

The mathematical model behind IRT transforms ordinal data into a linear and equal interval measure (Boone, Staver, & Yale, 2014). This new "person score" can be applied to parametric techniques such as unidimensional or multidimensional factor analyses without violating the assumptions typically held by CTT.

Establishing Measurement Equivalence

In this study, the method used to establish measurement invariance in the IRT framework was adapted from the work of Tay et al. (2015), in which they provide a comparison of the process to establish measurement equivalence (ME) from a CFA and IRT perspective. The primary form of analysis in the IRT framework is DIF (differential item functioning). It is important to note that, unlike traditional CFA-based measurement equivalence where the emphasis

is on comparing the construct level of groups, DIF focuses on the functioning of items across groups. This is not to suggest that test/scale score equivalence is irrelevant, as one can determine the extent to which DIF items cumulatively impact on observed mean differences across groups (i.e. differential test functioning [DTF]).

A bottom-up and progressively restricted methodology was adopted in this research to test for measurement invariance. In other words, the invariance requirements at each step needed to be met in order to move on to the higher-order form of equivalence (Dimitrov, 2010). The steps followed in order to determine ME are graphically depicted in Figure 1 and described below (Steenkamp & Baumgartner, 1998; Tay et al., 2015).

Configural Invariance

Configural or form invariance is the initial and most simple form of equivalence. In order to demonstrate configural invariance, the researcher needs to show that a factor is unidimensional across groups. Fundamentally, it is necessary to control for the underlying construct/latent trait (Cheung & Rensvold, 2002, Hair, Black, Babin, & Anderson, 2010).

Configural invariance was conducted to support local independence by establishing that observed item responses are uncorrelated after the following analyses:

- A restricted MIRT and/or unidimensional model (CFA equivalent) was conducted on each group within each DIF variable. In other words, the overall sample was split into two, three or four, depending on the number of subgroups. Separate models were conducted for annual turnover, and factor loadings were examined to assess if there were any significant differences.
- Thereafter, restricted multigroup MIRT modelling was conducted this analysed all groups within the same model to determine differences in terms of factor loadings and means. It is important to note that only items loading on designated factors were freely estimated similar to multifactor CFA analysis. This output helps to determining whether form equivalence exists for the total model of engagement.
- In order to identify specific areas of non-invariance, each subscale was analysed separately to examine differences that were flagged in the previous step. Unidimensional multigroup analyses also provided item-level statistics to assess the requirement of local independence.

The results of the above analyses were collated to assess whether the questionnaire data met the requirements of configural invariance. Thereafter, metric invariance was assessed via a DIF analysis. According to Hair et al. (2010), Dimitroy (2010), Raju et al. (2002) and Tay et al. (2015), metric invariance is a critical test of invariance, and the degree to which this is met, determines cross-group validity beyond the basic factor structure.

Figure 1. Measurement equivalence – IRT

Configural invariance

- Restricted MIRT
- Unidimentional IRT
- Restricted multigroup MIRT
- Multigroup unidimensional IRT

Metric Invariance

DIF unidimensional IRT

Scalar invariance

• Not tested as requirements for metric invariance not fully met.

RESEARCH PROCEDURE

The database of a research company was used in this study. The database in question consisted of approximately 285 000 businesspeople from various industries – including government institutions, sizes of business and who occupies different roles, reflecting the profile of the South African working population. The database was permissioned – that is, everybody in the database gave their permission to participate in online surveys. An electronic survey, administered by the iFeedback.co.za online data collection portal, was used by means of a mass e-mail invitation over a period of three weeks. Each potential participant received a personalised e-mail stating the purpose of the investigation and that the survey would take approximately 15 minutes to complete, and inviting them to participate in the survey on a voluntary, confidential and anonymous basis (Martins, 2015).

Population and Sampling

The majority of respondents resided in Gauteng (58.8%), the business hub of South Africa, and had been working at their company for over six years (74.3%).

Notably, 77.8% of the respondents were born before 1977, making them 38 years and older. In light of this information, it makes sense that many occupied executive, senior or management level positions at their organisation. A disproportionate percentage reported having a postgraduate qualification, while 54% also indicated that their organisation's annual turnover was more than R64 million. This implies that the sample consisted of a large number of formally educated and older individuals who held senior positions at companies.

The demographic items received few missing responses. Specific questions were asked about each industry to determine the exact sector of the respondents' organisations. These garnered the highest volume of missing responses.

Table 1. Breakdown of demographic variables

Item	Category	Frequency	Percentage
	0 to 1 year	166	4.0%
	2 to 3 years	456	11.1%
Years of service	4 to 5 years	429	10.5%
	6 to 10 years	1,006	24.5%
	10 years and longer	2,042	49.8%
	Born between 1946 and 1964	1,458	35.6%
Year of birth	Born between 1965 and 1977	1,730	42.2%
	Born between 1978 and 2000	911	22.2%
	Gauteng	2,408	58.8%
	Western Cape	754	18.4%
	KwaZulu-Natal	399	9.7%
	Free State	46	1.1%
D '	Limpopo	60	1.5%
Province	Mpumalanga	136	3.3%
	North West	91	2.2%
	Eastern Cape	152	3.7%
	Northern Cape	31	0.8%
	Not in South Africa	16	0.4%
C1	Male	2,387	58.2%
Gender	Female	1,712	41.8%
	Top management (Exco, COO, Director)	734	17.9%
<u> </u>	Senior management	1,149	28.1%
Job grade	Manager	1,175	28.7%
	Supervisor	349	8.5%
	Employee	688	16.8%
	Std 6 (Grade 8) and below	7	0.2%
	Stds 7 & 8 (Grades 9 & 10)	27	0.7%
	Stds 9 & 10 (Grades 11 & 12)	466	11.4%
Highest qualification	Certificate	495	12.1%
mgnest quantication	Diploma	952	23.2%
	First degree	745	18.2%
	Postgraduate qualification	1,404	34.3%
	Less than R0.20 million (R200 000)	77	1.9%
	Between R0.20 million and R6 million	0	0.0%
	Between R6 million and R32 million	427	10.4%
Annual turnover	Between R32 million and R64 million	321	7.8%
	More than R64 million	2,216	54.1%
	Do not know	1,058	25.8%

Research Questionnaire

The Employee Engagement Instrument (EEI), which was developed by Nienaber and Martins (2014) (and used with permission), measures engagement concurrently at the individual (organisational satisfaction and commitment), team (team and line manager) and organisational levels (strategy and customer service), in a multicultural environment. This instrument builds on existing research and was adapted for the South African context. Nienaber and Martins (2015: 16–17) reported on the validity and reliability of the instrument in the second phase of the research. The EEI questionnaire used for the current study consisted of eight biographical items, 50 Likert scale test items, the date of test completion and a unique identifier for each candidate. All identifying information was removed from the data to ensure ethical compliance and the respondents' anonymity. The sample size was 4 099 (Martins, 2015).

A Likert scale format was utilised to obtain information on employee engagement; the answers ranged along a fivepoint scale:

1 = strongly disagree; 2 = disagree; 3 = unsure; 4 = agree; 5 = strongly agree

The reliability results are reported below:

Table 2. Reliability statistics for the six extracted components

Description	Number of items	Cronbach's alpha
Team	11	0.933
Organisational Satisfaction	9	0.942
Immediate Manager	7	0.934
Organisational Commitment	12	0.932
Strategy and Implementation	7	0.904
Customer Service	4	0.813
All dimensions	50	0.973

Source: Martins & Ledimo (2016)

Analysis shows that the majority of respondents answered with options 3, 4 or 5. Options 1 and 2 were rarely selected by respondents. Items were negatively skewed with reported mean values of +- 3.5 to 4. Certain items received a broader range of responses, while others (e.g. those in the Team subscale) were more likely to generate positive answers of "agree" or "strongly agree".

No items reported missing data.

RESULTS

Unrestricted and Restricted Exploratory Factor Analysis (EFA)

EFA was conducted in SPSS 23 for the purposes of reviewing the underlying factor structure of the questionnaire. Both unrestricted (all items included) and restricted (only items for one subscale included) analyses were conducted. The method of extraction for both unrestricted and restricted EFAs was maximum likelihood with Promax Kappa 4 rotation. The number of extracted factors was based upon eigenvalues greater than 1. Factor loadings were sorted by size with coefficients smaller than 0.20 suppressed.

Kaiser-Meyer-Olkin's measure was performed and demonstrated sampling adequacy with all values greater than 0.86. Bartlett's test of sphericity was also conducted with p < 0.001 across all EFA analyses.

The restricted EFAs showed a single factor solution for each subscale, explaining between 47% and 67% of the total variance. Communalities were all above the accepted cut-off point of 0.50 (Hair et al., 2010).

The rotated pattern matrix for the unrestricted (all items included) EFA is depicted in Table 3. The item number and corresponding subscale are provided to give context to the factor loadings. Items whose factor loadings are highlighted in grey cross-load across two or more factors.

Table 3 Unrestricted EFA nattern matrix – CTT

	Table 3. Unrestricted EFA pattern matrix – CTT										
Item	Subscale				ctor						
no.		1	2	3	4	5	6				
Q41	Team	.904									
Q40	Team	.896									
Q39	Team	.808									
Q42	Team	.807									
Q36	Team	.764									
Q37	Team	.729									
Q50	Team	.708									
Q38	Team	.691									
Q35	Team	.686									
Q49	Team	.676									
Q51	Team	.674									
Q54	Team	.429					.311				
Q44	Customer Service	.323		.317							
Q11	Organisational Satisfaction		.972								
Q13	Organisational Satisfaction		.914								
Q12	Organisational Satisfaction		.870								
Q20	Organisational Satisfaction		.852								
Q10	Organisational Satisfaction		.838								
Q9	Organisational Satisfaction		.807								
Q23	Organisational Satisfaction		.730								
Q30	Organisational Satisfaction		.685								
Q34	Organisational Satisfaction	.200	.417	.243							
Q28	Strategy & Implementation			.963							
Q26	Strategy & Implementation			.949							
Q29	Strategy & Implementation			.771							
Q27	Strategy & Implementation			.757							
Q25	Strategy & Implementation			.719							
Q24	Strategy & Implementation			.602							
Q31	Strategy & Implementation			.599							
Q33	Customer Service	.286		.362							
Q32	Customer Service	.221		.328							
Q43	Customer Service	.221		.305							
Q55	Customer Service			.257							
Q33 Q47	Immediate Manager			.231	.964						
Q45	Immediate Manager				.926						
Q46	Immediate Manager				.925						
Q52	Immediate Manager				.778						
Q57	Immediate Manager				.685						
Q37 Q48	Immediate Manager				.659						
Q58	Immediate Manager			.231	.508						
Q38 Q16	Organisational Commitment	1		.431	.500	.852					
Q16 Q15	Organisational Commitment					.765					
	Organisational Commitment Organisational Commitment					.562					
Q17	Organisational Commitment Organisational Commitment					.520					
Q18							2.42				
Q19	Organisational Commitment		200			.406	.343				
Q14	Organisational Commitment	-	.288	274		.336	((2				
Q21	Strategy & Implementation			.274			.662				
Q22	Strategy & Implementation			222			.618				
Q53	Strategy & Implementation			.332			.602				
Q56	Customer Service			.226			.267				

Based on the results of the unrestricted EFA, six factors were identified and labelled as follows:

- **Factor 1:** Team (with one cross-loading item from Customer Service)
- Factor 2: Organisational Satisfaction
- Factor 3: Strategy and Implementation (with 4 items from Customer Service)
- Factor 4: Immediate Manager
- **Factor 5:** Organisational Commitment
- Factor 6: Mixed three items from Strategy and Implementation and one cross-loading item from Customer Service

It is evident that four out of the six factors were aligned to the theoretical factor structure of the questionnaire. However, items from Strategy and Implementation, and Customer Service did not load on to distinct factors; instead, they showed mixed results with items loading on non-designated factors. This is in line with the results of the restricted EFA that showed that Customer Service explained the lowest total variance out of the six subscales (47%).

Factor 6 reflects a number of items from Strategy and Implementation and Customer Service that centred around the strategy and future of the organisation. Item 56 ("My performance is linked to the strategy of the organisation") was theoretically clustered in Strategy and Implementation, and it thus made sense to move it to this dimension, owing to the fact that it loaded strongly on to said factor.

The items on factor 3 did not demonstrate any overt similarities. The group of items from Strategy and Implementation focused on accountability and encouragement. The group of items from Customer Service focused on going above and beyond for clients by providing feedback and exceeding expectations.

On factor 1, all items except one belonged to Team. The Customer Service item ("We identify the right opportunities for our customers") could have loaded here owing to the use of the term "we" - implying a team approach to servicing clients. However, this item was problematic in that it is cross-loaded on to both factor 1 and factor 3.

As discussed above, conducting an EFA within the CTT framework can be considered problematic, as ordinal data is treated as equal interval data. Also, Brown (2005) found that EFA is plagued by statistical artefacts such as overfactoring and Heywood cases, especially when a large number of items are factor analysed at the same time. The questionnaire data was thus re-examined from an IRT perspective. IRT analysis was conducted from both a unidimensional and multidimensional perspective – although multidimensional item response theory (MIRT) was the primary focus.

Unidimensional IRT analysis

All IRT analysis was conducted using IRTPRO version 3 for Windows. IRTPRO (item response theory for patientreported outcomes) models item calibration and scoring based on unidimensional and multidimensional versions of widely used IRT response functions.

For the unidimensional analyses, the Bock-Aitkin (BAEM) computational method (Bock & Aitkin, 1981) was utilised for item parameter estimation. The computation of scale scores was completed according to the expected a posteriori (EAP) method.

The results demonstrated that the IRT factor structure followed the same pattern evident in the restricted EFA. Table 4 provides a comparison of the six unidimensional models in terms of goodness of fit:

- -2loglikelihood is used as a relative comparison between competing models to assess fit. A smaller value is preferable. AIC and BIC are similarly interpreted.
- M_2 is a statistic that can be used as a proxy for chi-squared. Smaller values indicate a higher degree of model fit. A significant M_2 value (p < 0.05) is an indicator of model misfit.
- The root mean square error of approximation (RMSEA) is a calculation based on the chi-squared statistic. Values closer to 0 indicate good model fit (<0.05), while values closer to 1 indicate misfit.

From the results below, it is evident that all the subscales demonstrated some degree of model misfit.

Table 4. Goodness-of-fit comparison – unidimensional IRT

Subscale	-2loglikelihood	AIC	BIC	M ₂	RMSEA
Team	81 113.70	81 233.70	81 551.98	10 879.68***	0.05
Organisational Satisfaction	66 611	66 701	66 985.33	10 214.70***	0.06
Strategy & Implementation	93 213.9	93 313.9	93 629.83	91 74.78***	0.05
Customer Service	57 375.28	57 435.28	57 624.83	20 28.06***	0.04
Immediate Manager	60 455.77	60 525.77	60 746.92	67 71.07***	0.07
Organisational Commitment	53 351.86	53 411.86	53 601.42	3 931.18***	0.06

^{*** -} p < 0.001; AIC - Akaike information criterion; BIC - Bayesian information criterion

Unrestricted MIRT

The unrestricted multidimensional IRT (MIRT) analysis included all items of the scale-based questions and can be considered analogous to a CTT unrestricted EFA. For the purposes of the current study, it was compared to the outcomes obtained from the EFA to determine if the EEI's factor structure remained similar when analysed from an IRT perspective.

Item parameter estimation for the MIRT models throughout this study was calculated according to the Metropolis-Hastings Robbins-Monro (MHRM) method, recommended for analyses with more than two factors or dimensions (Cai, 2010).

The oblique rotated factor loadings for the unrestricted MIRT are depicted in Table 5. The values in brackets represent the factor loadings on the unrestricted EFA analysis as shown in Table 5. Those items highlighted cross-load on to multiple factors and could be deemed problematic.

Table 5. Unrestricted EFA and unrestricted MIRT factor loading comparison

		Cled EFA and	unrestricted MI	Fac		1	
Item	Subscale	1	2	3	4	5	6
Q45	Immediate Manager		0.96 (0.92)				
Q46	Immediate Manager		0.97 (0.92)				
Q47	Immediate Manager		0.97 (0.96)				
Q48	Immediate Manager		0.76 (0.66)				
Q52	Immediate Manager		0.83 (0.78)				
Q57	Immediate Manager		0.74 (0.68)				
Q58	Immediate Manager		0.58 (0.51)				
Q35	Team			-0.64 (0.69)			
Q36	Team			-0.73 (0.76)			
Q37	Team			-0.71 (0.73)			
Q38	Team			-0.71 (0.69)			
Q39	Team			-0.86 (0.81)			
Q40	Team			-0.92 (0.90)			
Q41	Team			-0.91 (0.90)			
Q42	Team			-0.85 (0.81)			
Q44	Customer Service			-0.38 (0.32)	0.24		0.25
Q49	Team			-0.68 (0.68)			
Q50	Team			-0.75 (0.71)			
Q51	Team			-0.69 (0.67)			
Q54	Team			-0.5 (0.43)			
Q55	Customer Service			-0.29	(0.26)		
Q43	Customer Service			-0.28	(0.30)		0.32
Q56	Customer Service	(0.27)		-0.28	,		0.31
Q24	Strategy and Implementation				0.52 (0.60)		
Q25	Strategy and Implementation				0.62 (0.72)		
Q26	Strategy and Implementation				0.74 (0.95)		
Q27	Strategy and Implementation				0.58 (0.76)		
Q28	Strategy and Implementation				0.73 (0.96)		
Q29	Strategy and Implementation				0.65 (0.77)		
Q31	Strategy and Implementation				0.46 (0.60)		
Q32	Customer Service			-0.28	0.3 (0.32)		
Q33	Customer Service			-0.33	0.28 (0.36)		
Q9	Organisational Satisfaction				()	0.88 (0.81)	
Q10	Organisational Satisfaction					0.91 (0.84)	
Q11	Organisational Satisfaction					0.94 (0.97)	
Q12	Organisational Satisfaction					0.88 (0.87)	
Q13	Organisational Satisfaction					0.92 (0.91)	
Q20	Organisational Satisfaction					0.85 (0.85)	
Q23	Organisational Satisfaction					0.78 (0.73)	
Q30	Organisational Satisfaction					0.74 (0.68)	
Q34	Organisational Satisfaction					0.42 (0.42)	
Q14	Organisational Commitment		 			(4.12)	0.42 (0.34)
Q15	Organisational Commitment						0.74 (0.76)
Q16	Organisational Commitment						0.79 (0.85)
Q17	Organisational Commitment						0.69 (0.56)
Q18	Organisational Commitment	1					0.69 (0.52)
Q19	Organisational Commitment						0.72 (0.41)
Q21	Strategy and Implementation	(0.66)	+				0.72 (0.41)
Q22	Strategy and Implementation	(0.62)					0.65
Q53	Strategy and Implementation	(0.60)					0.49

The metrics in Table 5 show a measure of correspondence between the EFA (CTT) results and the unrestricted MIRT. However, there were a few cases where factor loadings were not similar, or where items loaded on to different factors.

Restricted MIRT

Given the large degree of correspondence of the unrestricted MIRT to both the unrestricted EFA and the questionnaire's theoretical structure, a restricted MIRT was conducted that prespecified the factor loadings. In other words, all non-relevant factor loadings were constrained to zero, thus adopting a purely confirmatory approach.

The restricted MIRT can be considered the IRT analogue to the confirmatory factor analysis (CFA) conducted in Martins (2015). The results of the analysis are depicted in Table 6 below. Factor loadings from the CFA are added in brackets. Where an item was excluded from the CFA analysis from Martins (2015), it contains "N/A" in brackets.

In the case of the restricted MIRT, items were assigned to scales as defined by the theoretical structure of the EEI. The results show a similar factor structure between the original CFA and IRT approaches, with differences primarily arising in the case where CFA items were assigned to different factors.

Table 6. CFA and restricted MIRT factor loading comparison

T4	Chl-			Fac	tor		
Item	Subscale	1	2	3	4	5	6
Q32	Customer Service	0.65 (0.66)					
Q33	Customer Service	0.68 (0.69)					
Q43	Customer Service	0.76 (0.77)					
Q44	Customer Service	0.80 (0.78)					
Q55	Customer Service	0.65		(0.56)			
Q56	Customer Service	0.74		(0.64)			
Q45	Immediate Manager		0.92 (0.92)				
Q46	Immediate Manager		0.91 (0.89)				
Q47	Immediate Manager		0.91 (0.89)				
Q48	Immediate Manager		0.74 (0.65)				
Q52	Immediate Manager		0.86 (0.82)				
Q57	Immediate Manager		0.84 (0.80)				
Q58	Immediate Manager		0.78 (N/A)				
Q14	Organisational Commitment			0.77 (0.77)			
Q15	Organisational Commitment			0.86 (0.73)			
Q16	Organisational Commitment			0.90 (0.78)			
Q17	Organisational Commitment			0.80 (0.77)			
Q18	Organisational Commitment			0.81 (0.78)			
Q19	Organisational Commitment			0.85 (0.77)			
Q9	Organisational Satisfaction				0.84 (0.77)		
Q10	Organisational Satisfaction				0.85 (0.78)		
Q11	Organisational Satisfaction				0.91 (0.88)		
Q12	Organisational Satisfaction				0.91 (0.87)		
Q13	Organisational Satisfaction				0.90 (0.85)		
Q20	Organisational Satisfaction				0.89 (0.84)		
Q23	Organisational Satisfaction				0.82 (0.74)		
Q30	Organisational Satisfaction				0.86 (0.82)		
Q34	Organisational Satisfaction				0.68 (0.62)		
Q21	Strategy and Implementation			(0.78)		0.82	
Q22	Strategy and Implementation			(0.79)		0.82	
Q24	Strategy and Implementation					0.72 (0.66)	

(Table 5 continued on next page.

(Table 6 continued)

T4	Subscale			Fact	or		
Item	Subscale	1	2	3	4	5	6
Q25	Strategy and Implementation					0.64 (0.65)	
Q26	Strategy and Implementation					0.81 (0.78)	
Q27	Strategy and Implementation					0.85 (0.85)	
Q28	Strategy and Implementation					0.86 (0.82)	
Q29	Strategy and Implementation					0.64 (0.65)	
Q31	Strategy and Implementation					0.84 (0.84)	
Q53	Strategy and Implementation			(0.72)		0.77	
Q35	Team						0.77 (0.66)
Q36	Team						0.79 (0.67)
Q37	Team						0.80 (0.74)
Q38	Team						0.79 (0.73)
Q39	Team						0.83 (0.75)
Q40	Team						0.85 (0.74)
Q41	Team						0.86 (0.76)
Q42	Team						0.83 (0.77)
Q49	Team						0.84 (0.79)
Q50	Team						0.80 (0.75)
Q51	Team						0.79 (0.74)
Q54	Team			(0.37)			0.71

^{*1 -} Customer Service; 2 - Immediate Manager; 3 - Organisational Commitment; 4 - Organisational Satisfaction; 5 - Strategy & Implementation; 6 - Team

At this stage, the results of the analysis established a similar pattern of factor loadings across the EFA, CFA and IRT approaches. This provides strong evidence of the factor analytic structure of the measure as well as model fit for the individual scales of the EEI. Next, the focus of the report will shift to the measurement invariance of the EEI in relation to turnover

INVARIANCE TESTING: CONFIGURAL INVARIANCE

Most of the invariance analyses focused on invariance in specific personal characteristics such as gender, race, age and so forth, but seldom on a more business-related focus such as annual turnover.

A breakdown of the original response frequencies is provided in Table 1. The vast majority of respondents responded by selection a 5 or 6 for the turnover item. No respondents fell into the "Annual turnover more than R0.20 million but less than R6 million" category. For the purposes of the IRT analysis, "do not know" (6) was excluded, as these results were not relevant to the analysis. "Annual turnover less than R0.20 million" (1) and "Annual turnover more than R0.20 million but less than R6 million" (2) were excluded because of the relatively small sample sizes. Owing to the unbalanced size of the "Annual turnover more than R64 million" (5) responses, a randomised selection of 500 cases (Tay et al., 2015) was used for analysis. The remaining three categories were recoded as follows:

- 1. Annual turnover more than R6 million but less than R32 million (small)
- 2. Annual turnover more than R32 million but less than R64 million (medium)
- 3. Annual turnover more than 64 million (large)

The following steps were followed to investigate measurement invariance in relation to annual turnover:

- Unidimensional analyses were conducted on each group within the annual turnover variable to assess factor loadings and goodness of fit between said groups.
- A restricted multigroup MIRT was then conducted to examine the functioning of all subscales in the same model. This analysis included all four categories of annual turnover assessed simultaneously.

- Multigroup analyses were also conducted on a unidimensional level. This was done in order to obtain
 more detailed item level metrics. The multigroup unidimensional models provide the baseline group
 against which the DIF analysis can be compared.
- The above three steps contributed towards an assessment of configural invariance. The last step, namely the DIF analysis, was conducted to provide evidence towards determining metric invariance. A multigroup DIF analysis was run on each subscale to determine the item differences between groups. The results of each DIF were compared to the corresponding baseline model to determine the presence of DIF (or DTF) at a model level.

Per Group Unidimensional Analyses: Turnover

Table 7 provides a breakdown of goodness of fit of the unidimensional models run on each turnover group within each subscale. A total of 18 different models were analysed to account for the three categories within each of the six subscales.

Misfit was detected via the highly significant M₂ indicators. The second turnover response category ("Annual turnover more than R32 million but less than R64 million") reported a lower -2loglikelihood result, yet paradoxically a higher M₂ result, relative to the other groups in each subscale. The RMSEA reported fit indices that were more indicative of good model fit. Overall, no single group reported better fit relative to the other turnover categories. Since the M₂ statistic and -2loglikelihood are influenced by sample size, it is more prudent to place more weight on the RMSEA estimates (Hair et al., 2010). In this instance, a sample size of more than 500 was used. According to Tay et al. (2015), a sample size of more than 500 is suggested for IRT analysis.

Table 7. Goodness of fit comparison – turnover unidimensional analyses

Customer Service (6 items)									
Group	Sample size	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA		
More than R6 million but less than R32 million	427	5604.44	5664.44	5786.14	414.64***	234	0.04		
More than R32 million but less than R64 million	321	4247.48	4307.48	4420.62	469.19***	234	0.06		
More than R64 million	500	7073.56	7133.56	7260.00	529.28***	234	0.05		

Immediate Manager (7 items)										
Group	Sample size	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA			
More than R6 million but less than R32 million	427	6229.81	6299.81	6441.80	1196.03***	329	0.08			
More than R32 million but less than R64 million	321	4502.71	4572.71	4704.71	894.49***	329	0.07			
More than R64 million	500	7440.54	7510.54	7658.05	1927.13***	329	0.1			

Organisational Commitment (6 items)										
Group	Sample size	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA			
More than R6 million but less than R32 million	427	5465.03	5525.03	5646.73	758.7***	234	0.07			
More than R32 million but less than R64 million	321	4049.91	4109.91	4223.05	464.19***	234	0.06			
More than R64 million	500	6595.73	6655.73	6782.17	908.53***	234	0.08			

(Table 7 continued on next page)

(Table 7 continued)

Organisational Satisfaction (9 items)										
Group	Sample size	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA			
More than R6 million but	427	6937.4	7027.40	7209.96	1749.76***	567	0.07			
less than R32 million	427	0937.4	7027.40	1207.70	1747.70	307	0.07			
More than R32 million	321	4873.52	4963.52	5133.23	2995.87***	567	0.12			
but less than R64 million	321	48/3.32	4903.32	3133.23	2993.87	307	0.12			
More than R64 million	500	8214.52	8304.52	8494.18	2027.73***	567	0.07			

Strategy & Implementation (10 items)										
Group	Sample size	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA			
More than R6 million but less than R32 million	427	9624.11	9724.11	9926.95	1652.80***	710	0.06			
More than R32 million but less than R64 million	321	7010.67	7110.67	7299.24	2078.77***	710	0.08			
More than R64 million	500	11318.00	11418.00	11628.73	1557.59***	710	0.05			

Team (12 items)								
Group	Sample size	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA	
More than R6 million but less than R32 million	427	8279.56	8399.56	8642.97	2971.50***	1044	0.07	
More than R32 million but less than R64 million	321	5753.48	5873.48	6099.76	3013.99***	1044	0.08	
More than R64 million	500	9916.00	10036.00	10288.88	3118.63***	1044	0.06	

Restricted multigroup MIRT: Turnover

Within this restricted MIRT, all three turnover groups were included in one model, along with all 50 EEI items. For the purposes of analysis, turnover category 1 ("Annual turnover more than R6 million but less than R32 million") was treated as the reference group, while turnover groups 2 and 3 were categorised as focal groups. The choice of reference group was arbitrary.

The factor loadings of the restricted multigroup MIRT are represented in Tables 8 and 9. Group parameter estimates and latent variable variance-covariance matrices are then depicted in Tables 10 and 11.

Table 8. Multigroup restricted MIRT factor loadings – turnover group 1

		Toup restricted	WIIKT Tactor I	oadings – turn Fac			
Item	Subscale	1	2	3	4	5	6
Q32	Customer Service	0.73					
Q33	Customer Service	0.75					
Q43	Customer Service	0.81					
Q44	Customer Service	0.87					
Q55	Customer Service	0.74					
Q56	Customer Service	0.80					
Q45	Immediate Manager		0.96				
Q46	Immediate Manager		0.94				
Q47	Immediate Manager		0.93				
Q48	Immediate Manager		0.82				
Q52	Immediate Manager		0.85				
Q57	Immediate Manager		0.87				
Q58	Immediate Manager		0.81				
Q14	Organisational Commitment			0.86			
Q15	Organisational Commitment			0.90			
Q16	Organisational Commitment			0.96			
Q17	Organisational Commitment			0.79			
Q18	Organisational Commitment			0.84			
Q19	Organisational Commitment	İ		0.86			
Q9	Organisational Satisfaction			3,00	0.87		
Q10	Organisational Satisfaction				0.87		
Q11	Organisational Satisfaction				0.93		
Q12	Organisational Satisfaction				0.93		
Q13	Organisational Satisfaction				0.94		
Q20	Organisational Satisfaction				0.92		
Q23	Organisational Satisfaction				0.83		
Q30	Organisational Satisfaction				0.87		
Q34	Organisational Satisfaction				0.67		
Q21	Strategy & Implementation				0.07	0.87	
Q22	Strategy & Implementation					0.86	
Q24	Strategy & Implementation					0.77	
Q25	Strategy & Implementation					0.68	
Q26	Strategy & Implementation					0.88	
Q27	Strategy & Implementation					0.89	
Q28	Strategy & Implementation					0.91	
Q29	Strategy & Implementation					0.68	
Q31	Strategy & Implementation					0.86	
Q53	Strategy & Implementation					0.83	
Q35	Team					0.03	0.83
Q36	Team						0.82
Q37	Team						0.85
Q37 Q38	Team						0.86
Q39	Team						0.86
Q40	Team						0.89
Q40 Q41	Team						0.89
Q41 Q42	Team						0.85
Q42 Q49	Team						0.85
Q50	Team						0.88
Q51	Team						0.78
Q54	Team						0.83

Table 9. Multigroup restricted MIRT factor loadings – turnover group 2

	Table 9. Multig	Toup restricted	WIIKI Iactori	Fac			
Item	Subscale	1	2	3	4	5	6
Q32	Customer Service	0.62					
Q33	Customer Service	0.66					
Q43	Customer Service	0.72					
Q44	Customer Service	0.68					
Q55	Customer Service	0.62					
Q56	Customer Service	0.70					
Q45	Immediate Manager		0.69				
Q46	Immediate Manager		0.68				
Q47	Immediate Manager		0.67				
Q48	Immediate Manager		0.54				
Q52	Immediate Manager		0.65				
Q57	Immediate Manager		0.63				
Q58	Immediate Manager		0.59				
Q14	Organisational Commitment			0.59			
Q15	Organisational Commitment			0.63			
Q16	Organisational Commitment			0.67			
Q17	Organisational Commitment	İ		0.59			
Q18	Organisational Commitment	Ì		0.60			
Q19	Organisational Commitment	İ		0.64			
Q9	Organisational Satisfaction				0.73		
Q10	Organisational Satisfaction	İ			0.73		
Q11	Organisational Satisfaction	İ			0.77		
Q12	Organisational Satisfaction				0.78		
Q13	Organisational Satisfaction				0.80		
Q20	Organisational Satisfaction				0.77		
Q23	Organisational Satisfaction				0.73		
Q30	Organisational Satisfaction				0.75		
Q34	Organisational Satisfaction				0.62		
Q21	Strategy & Implementation				0.02	0.62	
Q22	Strategy & Implementation					0.62	
Q24	Strategy & Implementation					0.57	
Q25	Strategy & Implementation					0.53	
Q26	Strategy & Implementation					0.64	
Q27	Strategy & Implementation					0.66	
Q28	Strategy & Implementation	1				0.66	
Q29	Strategy & Implementation					0.5	
Q31	Strategy & Implementation					0.65	
Q53	Strategy & Implementation					0.64	
Q35	Team					0.01	0.82
Q36	Team						0.84
Q37	Team						0.86
Q37 Q38	Team	1					0.81
Q39	Team						0.85
Q40	Team	1					0.89
Q40 Q41	Team	1					0.87
Q41 Q42	Team	1					0.86
Q42 Q49	Team						0.84
Q49 Q50	Team	1					0.85
Q50 Q51	Team						0.85
Q51 Q54	Team	İ					0.83

Table 10. Multigroup restricted MIRT factor loadings – turnover group 3

	Table 10. Multigro	Jup restricted i	viiki iactoi t	Fac			
Item	Subscale	1	2	3	4	5	6
Q32	Customer Service	0.62					
Q33	Customer Service	0.61					
Q43	Customer Service	0.66					
Q44	Customer Service	0.69					
Q55	Customer Service	0.56					
Q56	Customer Service	0.67					
Q45	Immediate Manager		0.78				
Q46	Immediate Manager		0.79				
Q47	Immediate Manager		0.78				
Q48	Immediate Manager		0.65				
Q52	Immediate Manager		0.74				
Q57	Immediate Manager		0.71				
Q58	Immediate Manager		0.67				
Q14	Organisational Commitment			0.64			
Q15	Organisational Commitment			0.74			
Q16	Organisational Commitment			0.78			
Q17	Organisational Commitment			0.73			
Q18	Organisational Commitment			0.73			
Q19	Organisational Commitment			0.75			
Q9	Organisational Satisfaction			3175	0.75		
Q10	Organisational Satisfaction				0.76		
Q11	Organisational Satisfaction				0.81		
Q12	Organisational Satisfaction				0.81		
Q13	Organisational Satisfaction				0.79		
Q20	Organisational Satisfaction				0.79		
Q23	Organisational Satisfaction				0.73		
Q30	Organisational Satisfaction				0.78		
Q34	Organisational Satisfaction				0.63		
Q21	Strategy & Implementation				0.05	0.79	
Q22	Strategy & Implementation					0.8	
Q24	Strategy & Implementation					0.7	
Q25	Strategy & Implementation					0.67	
Q26	Strategy & Implementation					0.8	
Q27	Strategy & Implementation					0.84	
Q28	Strategy & Implementation					0.82	
Q29	Strategy & Implementation					0.65	
Q23	Strategy & Implementation Strategy & Implementation					0.81	
Q53	Strategy & Implementation					0.76	
Q35	Team					0.70	0.86
Q36	Team						0.84
Q37	Team						0.89
Q37 Q38	Team						0.84
Q38 Q39	Team						0.84
Q40	Team						0.91
Q40 Q41	Team						0.92
Q41 Q42	Team						0.98
Q42 Q49	Team						0.89
Q50	Team						0.93
Q50 Q51	Team						0.83
	1 Calli	1					0.83

In general, the factor loadings were strong across all three turnover groups. On average, the factor loadings were somewhat weaker for group 2. However, the factor structure replicated relatively well across the three groups. The latent mean estimates and standard errors are reported in Table 11. Since the reference group was "Annual turnover more than R6 million but less than R32 million", all the mean estimates across the six subscales seemed to be higher in the other two groups. It is important to reiterate that group mean comparisons are only permissible if the assumption of metric invariance holds across the majority of items in the subscales. According to this assumption, it would seem that the candidates in group 2 differed most from group 1.

Table 11. Multigroup restricted MIRT group parameter estimates – turnover

Label	Group	μ_1	s.e.	μ_2	s.e.	μ3	s.e.	μ_4	s.e.	μ ₅	s.e.	μ_6	s.e.
1	More than R6 million but less than R32 million	0		0		0		0		0		0	
2	More than R32 million but less than R64 million	0.71	0.08	0.46	0.11	0.69	0.10	0.64	0.08	0.34	0.09	0.76	0.06
3	More than R64 million	0.33	0.06	0.35	0.05	0.36	0.05	0.56	0.05	0.16	0.05	0.73	0.03

The latent covariances are reported in Table 11. The covariances between the latent variables were all positive as would be expected. None of the standard errors were abnormally high.

Table 12. Multigroup restricted MIRT latent variable variance-covariance matrix – turnover

	Group 1: More than R6 million but less than R32 million										
θ_1	s.e.	θ2	s.e.	θ3	s.e.	θ4	s.e.	θ5	s.e.	θ6	s.e.
1											
0		1									
0		0		1							
0		0		0		1					
0		0		0		0		1			
0		0		0		0		0		1	

	Group 2: More than R32 million but less than R64 million										
θ_1	s.e.	θ2	s.e.	θ3	s.e.	θ4	s.e.	θ5	s.e.	θ6	s.e.
1.38	0.20										
1.22	0.17	1.92	0.18								
1.32	0.17	1.29	0.16	1.92	0.17						
1.10	0.12	1.1	0.12	1.37	0.13	1.41	0.13				
1.43	0.22	1.55	0.19	1.59	0.16	1.21	0.13	1.88	0.27		
0.89	0.14	0.86	0.11	0.71	0.15	0.76	0.10	0.81	0.16	1.05	0.11

	Group 3: More than R64 million										
θ_1	s.e.	θ2	s.e.	θ3	s.e.	θ4	s.e.	θ5	s.e.	θ6	s.e.
1.38	0.08										
0.94	0.05	1.43	0.06								
1.16	0.07	0.93	0.07	1.39	0.08						
1.00	0.07	0.88	0.06	1.12	0.05	1.30	0.08				
1.07	0.07	0.91	0.07	1.08	0.08	0.86	0.07	1.14	0.08		
0.87	0.06	0.61		0.63	0.05	0.67	0.05	0.59	0.06	0.88	0.06

Unidimensional Multigroup Analyses (Baseline Models) – Turnover

At this point, each subscale was analysed with all three turnover groups being simultaneously estimated in each model. As previously mentioned, these baseline analyses provide an indication of configural invariance as well as a comparison point for subsequent analyses of invariance (Hair et al., 2010).

Table 13 presents the factor loadings for each baseline multigroup analysis. The results in this table suggest a high degree of consistency in the magnitude of factor loadings across the three turnover groups. However, the team subscale was problematic across the three groups and some abnormal factor loadings were reported.

Table 13. Multigroup baseline factor loadings – turnover

	Customer Service								
Item	6-32 million (1)	32-64 million (2)	More than 64 million (3)						
Q32	0.71	0.81	0.83						
Q33	0.74	0.85	0.81						
Q43	0.82	0.85	0.82						
Q44	0.87	0.86	0.94						
Q55	0.75	0.77	0.71						
Q56	0.80	0.86	0.78						

	Immediate Manager								
Item	6-32 million (1)	32-64 million (2)	More than 64 million (3)						
Q45	0.96	0.88	0.88						
Q46	0.94	0.87	0.89						
Q47	0.93	0.86	0.88						
Q48	0.83	0.68	0.72						
Q52	0.85	0.81	0.82						
Q57	0.87	0.79	0.79						
Q58	0.82	0.74	0.73						

	Organisational Commitment								
Item	6-32 million (1)	32-64 million (2)	More than 64 million (3)						
Q14	0.87	0.73	0.70						
Q15	0.91	0.81	0.84						
Q16	0.96	0.84	0.88						
Q17	0.79	0.73	0.81						
Q18	0.85	0.75	0.79						
Q19	0.87	0.80	0.81						

Organisational Satisfaction								
Item	6-32 million (1)	32-64 million (2)	More than 64 million (3)					
Q9	0.88	0.89	0.89					
Q10	0.88	0.88	0.91					
Q11	0.93	0.94	0.97					
Q12	0.93	0.95	0.95					
Q13	0.94	0.97	0.94					
Q20	0.93	0.94	0.94					
Q23	0.84	0.88	0.87					
Q30	0.88	0.90	0.92					
Q34	0.67	0.74	0.74					

(Table 13 continued on next page)

(Table 13 continued)

	(14010 13	continucaj									
	Strategy & Implementation										
Item	6-32 million (1)	32-64 million (2)	More than 64 million (3)								
Q21	0.87	0.83	0.90								
Q22	0.86	0.83	0.89								
Q24	0.78	0.80	0.81								
Q25	0.70	0.75	0.80								
Q26	0.89	0.90	0.93								
Q27	0.90	0.94	0.96								
Q28	0.91	0.92	0.94								
Q29	0.69	0.72	0.76								
Q31	0.86	0.89	0.91								
Q53	0.83	0.87	0.86								

	Team										
Item	6-32 million (1)	32-64 million (2)	More than 64 million (3)								
Q35	0.82	0.95	0.93								
Q36	0.82	0.98	0.91								
Q37	0.84	1.01	0.96								
Q38	0.86	0.95	0.92								
Q39	0.86	1.00	0.99								
Q40	0.89	1.04	1.02								
Q41	0.88	1.01	1.06								
Q42	0.85	1.00	0.96								
Q49	0.86	0.99	1.00								
Q50	0.78	1.00	0.93								
Q51	0.82	0.99	0.89								
Q54	0.72	0.84	0.82								

In Table 14, the latent means, variance and standard deviation are reported with associated standard error estimates.

Table 14. Multigroup parameter estimates – turnover

Custome	Customer Service											
Label	Group	M	s.e.	σ^2	s.e.	σ	s.e.					
1	6-32 million	0		1		1						
2	32-64 million	0.43	0.12	0.89	0.30	0.94	0.16					
3	More than 64 million	0.35	0.12	0.82	0.73	0.91	0.40					

Immediate Manager										
Label	Group	M	s.e.	σ^2	s.e.	σ	s.e.			
1	6-32 million	0		1		1				
2	32-64 million	0.34	0.08	1.20	0.22	1.09	0.10			
3	More than 64 million	0.24	0.06	1.13	0.19	1.06	0.09			

	Organisational Commitment										
Label	Group	M	s.e.	σ^2	s.e.	σ	s.e.				
1	6-32 million	0		1		1					
2	32-64 million	0.56	0.11	1.23	0.32	1.11	0.14				
3	More than 64 million	0.41	0.09	1.12	0.26	1.06	0.12				

	Organisational Satisfaction										
Label	Label Group M s.e. σ ² s.e. σ s.e.										
1	6-32 million	0		1		1					
2	32-64 million	0.60	0.10	0.96	0.21	0.98	0.11				
3	More than 64 million	0.49	0.05	0.91	0.10	0.96	0.05				

	Strategy & Implementation										
Label	Group	M	s.e.	σ^2	s.e.	σ	s.e.				
1	6-32 million	0		1		1					
2	32-64 million	0.25	0.07	0.98	0.20	0.99	0.10				
3	More than 64 million	0.19	0.23	0.87	2.67	0.93	1.44				

	Team										
Label	Group	M	s.e.	σ^2	s.e.	σ	s.e.				
1	6-32 million	0		1		1					
2	32-64 million	0.75	0.11	0.77	0.13	0.88	0.07				
3	More than 64 million	0.67	0.07	0.75	0.21	0.86	0.12				

The fit indices in Table 15 relate to the multigroup unidimensional models and are used as a reference or point of comparison to examine the assumption of metric invariance on the subscale level. Since DIF analyses take place on the subscale level, it does not make sense to compare the overall fit of the six subscales with one another. These estimates merely serve are baseline models to compare progressively restricted models of measurement invariance.

Table 15. Multigroup model fit – turnover

Model	-2loglikelihood	AIC	BIC	M ₂	d.f.	RMSEA
Customer Service	16925.48	17113.48	17595.63	1394.55***	698	0.03
Immediate Manager	18173.06	18391.06	18950.15	4008.87***	983	0.05
Organisational Commitment	16110.67	16298.67	16780.82	1795.51***	698	0.04
Organisational Satisfaction	20025.43	20303.43	21016.40	6700.96***	1697	0.05
Strategy & Implementation	27952.78	28260.78	29050.69	5143.88***	2126	0.03
Team	23949.03	24317.03	25260.82	9003.53***	3128	0.04

DIF Analysis - Turnover

Wald's χ^2 statistic implies significant differences between groups. In the three group analyses, contrast 1 compares group 1 with groups 2 and 3 (1 vs 2, 3). This can be regarded as the overall omnibus significance test similar to ANOVA. Contrast 2 compares groups 2 and 3 (2 vs 3).

Conducting post hoc analyses on annual turnover pairs (with group 1 as the consistent reference group) found evidence of limited DIF.

Overall, only one item on the Customer Service subscale demonstrated DIF according to the overall test (contrast 1). DIF post hoc analyses found no evidence of DIF between group 1 ("More than R6 million but less than R32 million") and group 2 ("More than R32 million but less than R64 million"). There was, however, one item with highly significant DIF between group 1 and group 3 ("More than R64 million").

```
• Group 1 and Group 3
Q33 (Total χ2 = 21.5; p<0.001)
```

Overall, only one item on the Immediate Manager subscale demonstrated DIF according to the overall test (contrast 1). When consulting the paired group analyses, the item differences were found to lie between Groups 1 and 3.

```
• Group 1 and Group 3
Q58 (Total \chi^2 = 17.1; p<0.01)
```

Both the multigroup and paired DIF analyses found no evidence of item differences between groups on Organisational Commitment. Strong support was found for metric invariance for the Organisational Commitment subscale. Hence the group mean scores could be compared on this subscale with confidence.

Overall, only one item on the Organisational Satisfaction subscale demonstrated DIF according to the overall test (contrast 1). Item 23 demonstrated DIF across groups 1 and 2, as well as across groups 1 and 3.

```
• Group 1 and Group 2
Q23 (Total \chi^2 = 11.7; p<0.05)
```

```
• Group 1 and Group 3
Q23 (Total \chi^2 = 14.3; p<0.05)
```

No DIF was found for the Strategy and Implementation subscale. Both the multigroup and paired DIF analyses found no evidence of item differences between the groups on Strategy and Implementation. Strong support was found for metric invariance for the subscale. Hence group mean scores could be compared with confidence.

Overall, only one item in the Team subscale demonstrated DIF according to the overall test (contrast 1). Post hoc analyses of DIF for the Team subscale found the following evidence of differential functioning:

```
• Group 1 and Group 2
Q 49 (Total \chi^2 = 16.2; p < 0.01)
```

```
• Group 1 and Group 3
Q38 (Total \chi^2 = 12.3; p < 0.05)
Q49 (Total \chi^2 = 12.7; p < 0.05)
```

In Table 16, the baseline models of each of the dimensions are compared to the DIF model to determine metric invariance.

Table 16. Comparison between Bit and baseine models turnover									
Analysis	-2loglikelihood	AIC	BIC	M2	d.f.	RMSEA	χ^2 Diff.		
Customer Service – Baseline	16925.48	17113.48	17595.63	1394.55***	698	0.03	56.66***		
Customer Service – DIF	16982.14	17132.14	17516.84	1482.70***	717	0.03	30.00		
Immediate Manager – Baseline	18173.06	18391.06	18950.15	4008.87***	983	0.05	352.84***		
Immediate Manager – DIF	18525.90	18699.90	19146.15	4521.44***	1005	0.05	332.84		
Organisational Commitment - Baseline	16110.67	16298.67	16780.82	1795.51***	698	0.04	157.04***		
Organisational Commitment - DIF	16267.71	16417.71	16802.41	2019.54***	717	0.04	137.04		
Organisational Satisfaction – Baseline	20025.43	20303.43	21016.40	6700.96***	1697	0.05	373.67***		
Organisational Satisfaction – DIF	20399.10	20621.10	21190.45	7172.82***	1725	0.05	3/3.0/		
Strategy & Implementation – Baseline	27952.78	28260.78	29050.69	5143.88***	2126	0.03	350.37***		
Strategy & Implementation – DIF	28303.15	28549.15	29180.05	6042.03***	2157	0.04	330.37		
Team – Baseline	23949.03	24317.03	25260.82	9003.53***	3128	0.04	164.78***		
Team DIE	2/11/2 91	24407.81	25161.92	0227 84***	2165	0.04	104./8		

Table 16. Comparison between DIF and baseline models – turnover

Team - DIF

The results in Table 16 indicate a significant deterioration in the -2loglikelihood metric when slopes are constrained equal across groups. All x^2 Diff. indicate significant differences between the baseline and restricted DIF models. With due consideration of the results reported in Table 16, it would be imprudent to compare the latent means of subscales directly across the three groups, since the assumption of metric invariance was partially satisfied.

24407.81

24113.81

25161.82

9237.84

3165

0.04

CONCLUSION

The main objectives of this study were as follows: (1) to confirm the factor analytic structure of the measure; (2) to determine the model fit for the individual scales of the EEI; and (3) to determine the measurement invariance of the EEI as a function of annual turnover.

Previous results have successfully demonstrated evidence of configural invariance using a CFA approach. The current study undertook to assess the three steps of measurement invariance according to the IRT framework. The measurement assumptions and properties of IRT frameworks are superior to classic test theory approaches.

The IRT analyses found support for the assumption of configural invariance of the total measure as well as for the subscales. Factor loadings were robust, with only minimal differences across groups and acceptable standard error values. The first objective was thus achieved.

However, with regard to the second objective, these results were tempered by relatively poor model fit and violation of the local independence requirement of IRT for some of the dimensions. This violation was most pronounced for the team subscale. It is posited that certain problematic items lead to a source of error in the analysis that has not been accounted for, such as multidimensionality.

Unidimensional DIF analyses were conducted at a subscale level to investigate metric invariance. The item slopes for each subscale were constrained equal, and the fit of the nested models was compared.

The -2loglikelihood difference test demonstrated a statistically significant difference in the model fit of between the baseline and restricted DIF models. This violates the requirements of metric invariance, namely that the slopes of the response categories in the graded model are similar across groups. Metric invariance can be considered an important prerequisite for comparing latent means between groups. Thus, the analyses demonstrated that for the most part it is not permissible to compare latent means directly across groups, especially for the Team dimension. In order to remedy this problem, items demonstrating DIF should be removed or adapted and retested.

With regard to the third objective, namely to determine the measurement of the EEI as a function of annual turnover, limited evidence of DIF was found. The results indicated only a few items demonstrating statistically different functioning as a result of annual turnover. Overall, only seven items demonstrated DIF, indicating, firstly, that the results of these items need to be interpreted with caution if the questionnaire is to be used in its current format. Most

^{*** -} p < 0.001; χ^2 Diff. – chi-squared difference test

of the items can thus be used with confidence in future analysis pertaining to annual turnover. However, it is suggested that the seven mentioned items should be adapted or removed.

Another suggestion is that the model fit for the individual scales is conducted for other biographical variables such as race and gender.

As mentioned earlier, one of the limitations of the research was the sample population which focused well-educated and older individuals with senior positions at well-performing companies. It would be of value for future research to include more respondents at all job and educational levels.

AUTHOR BIOGRAPHIES

Nico Martins. Articles which Nico Martins has published have dealt with organisational culture and climate, organisational diagnoses, employment equity and organisational trust. He has presented papers at more than 50 national and international conferences based on work done at various national and international companies. He has participated in about 500 reports on various aspects of organisational diagnoses, and has extensive knowledge in the fields of organisational assessment, (re)design and interventions. His research focuses on quantitative and qualitative research. E-mail: martin@unisa.ac.za (contact author)

REFERENCES

- Abrams, M. D., Lee, I. H., Brown, S. D., & Carr, A. (2015). The career indecision profile: Measurement equivalence in the United States and South Korea. Journal of Career Assessment, 23(2), 225–235.
- Bock, R. D., & Aitkin, M. (1981). Marginal maximum likelihood estimation of item parameters: An application of the EM algorithm. Psychometrika, 46, 443-459.
- Bond, T. G., & Fox, C. M. (2007). Applying the Rasch model: Fundamental measurement in the human sciences (2nd ed.). Mahwah, NJ: Erlbaum.
- Boone, W.J., Staver, J. R., & Yale, M. S. (2014), Rasch analysis in the human sciences Dordrecht: Springer.
- Brown, J. D. (2005). Testing in language programs: A comprehensive guide to English language assessment. New York:
- Byrne, B. M. (2004). Testing for multigroup invariance using AMOS graphics: A road less traveled. Structural Equation Modeling: A Multidisciplinary Journal, 11(2), 272–300. doi.org/10.1207/s15328007sem1102 8.
- Cai, L. (2010). High dimensional exploratory item factor analysis by a Metropolis-Hastings Robbins-Monro algorithm", Psychometrika, 1(75),33-57.
- Cheung, G. W. & Rensvold, R.B. (2002). Evaluating goodness-of-fit indexes for testing measurement invariance. Structural Equation Modeling, 9(2), 233-255.
- Dawson, D. B., Sotelo, F. L., Roesch, S. C., & Klonoff, E. A. (2014). Testing the invariance of adolescent survey-based smokingrelated behaviors across ethnic groups and gender. Journal of Health Psychology, 19(12), 1576–1585.
- Dimitrov, D. M. (2010). Testing for factorial invariance in the context of construct validation. Measurement and Evaluation in Counseling and Development, 43(2) 121–149.
- Egri, C. P. (2013). From the editors: Context matters in management education scholarship. Academy of Management Learning & Education, 12(2), 155–157. DOI.org/10.5465/amle.2013.0140.
- Embretson, S. E., & Reise, S. P. (2000). Item response theory for psychologists. Mahwah, NJ: Erlbaum.
- Fahr, J. L., Cannella, A. A., & Lee C. (2006). Approaches to scale development in Chinese management research. Management and Organization Review, 2(3), 301-318. DOI: 10.1111/j.1740-8784.2006.0055x.
- Fan, X. (1998). Item response theory and classical test theory: An empirical comparison of their item/person parameters. Educational and Psychological Measurement, 58, 357–381.
- González, D. A., & Jenkins, S. R. (2014). Cross-measure equivalence and communicability in the assessment of depression: A focus on factor-based sScales. Assessment, 21(6), 731–741.
- Hair, J. F., Black, W. C., Babin B. J., & Anderson, R. E. (2010). Multivariate analysis: A global perspective. (7th ed.). London.
- Harter J. K., Schmidt, F. L., & Hayes, T. L. (2002). Business-unit-level relationship between employee satisfaction, employee engagement, and business outcomes: A meta-analysis. Journal of Applied Psychology 87(2), 268–279.
- Hays, R. D., Morales, L. S., & Reise, S. P. (2000). Item response theory and health outcomes measurement in the 21st century. Medical Care, 38(9 Suppl.), II28-II42.
- Ho Yu, C. (2010). A simple guide to the item response theory (IRT) and Rasch modelling. Creative wisdom. Retrieved from http://www.creative-wisdom.com/.

- Horn, J. L., & McArdle, J. J. (1992). A practical and theoretical guide to measurement invariance in aging research. Experimental Aging Research, 18, 117-144.
- Linacre, J.M. (2005). Correlation coefficients: Describing relationships. Rasch measurement transactions, 19(3), 1028–1029. Retrieved from http://www.rasch.org/rmt/ rmt193c.htm
- Imandin, L., Bisschoff, C., & Botha, C. (2014). A model to measure employee engagement. Problems and Perspectives in Management, 12(4), 520-532.
- Johns, G. (2006). The essential impact of context on organizational behavior. Academy of Management Review, 31(2), 386-408.
- La Greca, A. M., Ingles, G. J., Lai, B. S., & Marco, J. C. (2015). Social anxiety scale for adolescents: Factorial invariance across gender and age in Hispanic American adolescents. Assessment, 22(2) 224–232.
- Lord, F. M. (1980). Applications of item response theory to practical testing problems. Hillsdale, NJ: Erlbaum.
- Makransky, G., & Bilenberg, N. (2014). Psychometric properties of the parent and teacher ADHD rating scale (ADHD-RS): Measurement invariance across gender, age, and informant. Assessment, 21(6), 694–705.
- Martins, N. (2014). Factorial invariance of the South African culture instrument. Problems and Perspectives in Management, 12(4), 242–252.
- Martins, N., & Da Veiga, A. (2015). Factorial invariance of an information security culture assessment instrument for multinational organisations with operations across data protection jurisdictions. Journal of Governance and Regulation,
- Martins, N. & Ledimo, O. (2016). The measurement of employee engagement in government institutions. *Journal Risk* Governance and Control, 6(3), 18-25.
- Meade, A.W., & Lautenschlarer, G.J. (2004). A comparison of Item Response Theory and Confirmatory Factor Analytic methodologies for establishing measurement Equivalence/Invariance, Organizational Research Methods, 7, 361-388.
- Meiring, D., Van de Vijver, A. J. R., Rothmann, S., & Barrick, M. R. (2005). Construct, item, and method bias of cognitive and personality tests in South Africa. SA Journal of Industrial Psychology, 31(1), 1-8. http//dx.doi.org/10.4102/sajip.v31i1.182.
- Milfont, T. L., & Fischer, R. (2010). Testing measurement invariance across groups: Applications in cross-cultural research. *International Journal of Psychological Research*, 3(1), 111–121.
- Moerdyk, A. (2009). The principles and practice of psychological assessment. Pretoria: Van Schaik.
- Nel, P., & Boshof, A. (2014). Factorial invariance of the Adult State Hope Scale. SA Journal of Industrial Psychology/SA Tydskrif vir Bedryfsielkunde, 40(1), Art.#1177, 8 pages.http//dx.doi.org/10.4102/sajip.v40i1.1177.
- Nienaber, H., & Martins, N. (2015). Validating a scale measuring engagement in a South African context. Journal of Contemporary Management, 12, 401–425.
- Raju, N. S., Laffitte, L. J., & Byrne, B. M. (2002). Measurement equivalence: A comparison of methods based on confirmatory factor analysis and item response theory. Journal of Applied Psychology, 87(3), 517–529.
- Reio, T. G., & Shuck, B. (2015). Exploratory factor analysis: Implications for theory, research, and practice. Advances in Developing Human Resources, 17(1), 12-25.
- Reise, S. P., Widaman, K. F., & Pugh, R. H. (1993). Confirmatory factora Analysis and item response theory: Two approaches for exploring measurement invariance. Psychological Bulletin, 114(3), 552-566.
- Republic of South Africa. (1996). National Small Business Act 102 of 1996. Pretoria: Government Printer. [http://www. acts.co.za/...act-1996/national small business...]
- Robbinson, D. & Gifford, J. (2014). The future of engagement thought piece collection. Brighton, UK: Institute for Employment
- Rock, D. A., Werts, C. E., & Flaugher, R. L. (1978). The use of analysis of covariance structures for comparing the osychometric properties of multiple variables across populations. Multivariate Behavioral Research, 13, 4, 403–418. DOI: 10.1207/s15327906mbr1304 3.
- Rouseau, D. M., & Fried, Y. (2001). Location, location; contextualizing organizational research. Journal of Organizational Behavior, 22, 1-13.
- Samuel, D. B., South, S. C., & Griffin, A. G. (2015). Factorial invariance of the five-factor model rating form across gender. Assessment, 22(1), 65-75.
- Spurk, D., Abele, A. E., & Volmer, J. (2015). The Career Satisfaction Scale in context: A test for measurement invariance across four occupational groups. Journal of Career Assessment, 23(2), 191-209.
- Strasheim, A. (2011) Testing the invariance of second-order confirmatory factor analysis models that include means and intercepts. Management Dynamics, 20(4), 38-75.
- Steenkamp, J-B. E. M., & Baumgartner, H. (1998). Assessing measurement invariance in cross-national consumer research. Journal of Consumer Research, 5, 78–90.
- Tay, L., Meade, A. W., & Cao, M. (2015). An overview and practical guide to IRT measurement equivalence analysis. Organizational Research Methods, 18(1), 3-46.
- Tsui, A. S. (2006). Contributing to global management knowledge: A case for high quality indigenous research. Asia Pacific Journal of Management, 21, 491–533.
- Vandenberg, R. J., & Lance, C. E. (2000). A review and synthesis of the measurement invariance literature: Suggestions,

practices, and recommendations for organizational research. *Organizational Research Methods*, *3*, 4–69. Whetten, D. A. (2009). An examination of the interface between context and theory applied to the study of Chinese organizations. *Management and Organization Review*, 5(1), 29–55. DOI: 10.1111/j.1740-8784.2008.00132.x.

NOTES