Using Multilevel Confirmatory Factor Analysis To Study Student Well-Being In Thailand

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ABSTRACT

Multilevel confirmatory factor analysis was used to assess the factor structure of the six-factor student well-being scale at between- and within-class levels. A total of 2,707 Matthayom 4-6 students (equivalent to grades 10-12) from 71 classrooms (785 male and 1,922 female) completed this 54-items student well-being scale. Results showed that a single latent factor structure was acceptable fitted at between and within class levels. The reliability estimates of the six factors (Positive attitudes and emotions towards school, Enjoyment in school, Positive academic self-concept, (Absence of) Social problems in school, (Absence of) School worries, and (Absence of) Physical complaints) were 0.73, 0.38, 0.22, 0.13, 0.14, and 0.02, respectively, at the student level, whereas these estimates were 0.39, 0.44, 0.48, 0.61, 0.65, and 0.82, respectively, at the classroom level.

Keywords: Student Well-Being; Multilevel Structural Equation Modeling; Multilevel Confirmatory Factor Analysis for Thai Students

INTRODUCTION

In Thailand, the importance of considering student well-being while forming education policies was recently pointed out. In accordance with national policy, the quality of students’ performance should be determined by not only cognitive but also affective criteria. Thus, Thai students are expected to excel academically, maintain good conduct in school, and be satisfied with their school lives. Although many research tools (e.g., standardized tests and achievement tests) are available to assess how and to what extent students perform academically, very little is known about the concept of student happiness or, more academically, “student well-being”. Borrowed from the psychological concept of well-being, this concept provides a clear and concrete basis for this study. Student well-being is a construct that describes an individual’s perception with regard to self-assessment of his/her life. It is defined as “an emotional experience characterized by the dominance of positive feeling and cognitions towards school, persons in school and the school context in comparison to negative feelings and cognitions towards school life. Well-being in school represents subjective, emotional and cognitive evaluations of school reality and can be seen as a misbalance of positive and negative aspects in favor of positive aspects” (Hascher, 2003, p.129 as cited in Hascher, 2008). Within the educational context, student well-being can create a positive learning environment that enables students to succeed academically and socially in school. Although it is not a direct school outcome in the way an achievement is, student well-being is required to create a positive bias for the learning experience in an education setting.

To examine the structure of the concept of student well-being, researchers usually employ confirmatory factor analysis for an entire sample (i.e., they pool all respondents from all classrooms), ignoring class-level variances in the factor structure. However, the conventional confirmatory factor analysis with its assumption of independent and identically distributed observations cannot be applied to this pooled sample, because students are clustered in the classroom. Thus, the method appropriate for this study is multilevel confirmatory factor analysis.
(MCFA), which fits the factor structure at both individual and class levels simultaneously and, further, gives lower bias for parameter estimation. Other advantages of this method include robust estimation with even unbalanced cases (Muthén, 1991, 1994; Hox & Maas, 2001).

Because of the complexity of multilevel data, guidelines to simplify the multilevel modeling process are needed. Muthén (1994) suggests conducting four initial steps before applying MCFA: applying a conventional factor analysis on the sample’s total covariance matrix ($S_T$), estimation of between variation, estimation of within structure ($S_W$), and estimation of between structures ($S_B$).

The advantages of Muthén’s procedures are that they allow:

- examination of various aspects of student well-being structures in single-level modeling, two-level modeling while ignoring clustered data, and two-level modeling with clustered data;
- exploration of the possible ideal of the modeling; and
- reduction of the technical problems regarding convergence, particularly the non-positive definite ones.

The two research questions in this study are:

- Is the factor structure underlining student well-being at the student level similar to or different from the factor structure at the class level?
- What is the reliability of the scores of student well-being at student and class levels?

**METHOD**

**Sample**

A sample of 2,707 Matthayom 4-6 students (equivalent to Grade 10-12) attending public schools run by the Office of Basic Education Commissions, Ministry of Education, Thailand participated in this study. A multi-stage sampling strategy based on regional and provincial criteria was used to select them. A sample of 36 schools was drawn from a database and 71 classes were selected from among them. Finally, data of all 2,707 students in those classes were used to perform the analysis.

**Questionnaire on student well-being**

A self-evaluation questionnaire developed by Hascher (2008) to examine student well-being within the school context was used. The questionnaire used a five-point Likert scale, ranging from 1 “strongly disagree” to 5 “strongly agree,” and assessed the following three positive and three negative dimensions: Positive attitudes and emotions towards school (9 items), for example, “I like to go to school”; Enjoyment in school (9 items), for example, “I have fun with the activities in class”; Positive academic self-concept (9 items), for example, “I like to share my opinions when teachers ask for it”; Social problems in school (9 items), for example, “I do not need friends in school”; School worries (9 items), for example, “Have you been worried in the past few weeks because of people whom you are dealing with”; Physical complaints (9 items), for example, “In the past few weeks, have you been to the nursing office?” The questionnaire was written in Thai. It was reviewed and approved by experts in the field of education and psychology before being administered.

**Data analysis**

The data analysis of student well-being was conducted using SPSS for descriptive information and Mplus Version 5.21 for model validity. The goodness of fit for the model was assessed using a chi-square statistic, comparative fit index (CFI), root mean square error of approximation (RMSEA), and standardized root mean square residual (SRMR). Acceptable fit was judged according to the criteria recommended by Hu and Bentler (1999): CFI and TLI values greater than or equal to .95 and RMSEA and SRMR values less than or equal to .06 and .08, respectively.
RESULTS

Descriptive Statistics

Item means ranged from 3.421 (SD = 0.610) for Social problems in school to 4.212 (SD = 0.506) for Positive attitudes and emotions towards school. Responses were assumed to be normally distributed, with skewness ranging from -0.736 to 0.023 and kurtosis values ranging from -0.059 to 0.832 (Table 1).

Multilevel Confirmatory Factor Analysis (MCFA)

Because of the complexity of the multilevel structure, the literature recommends conducting the aforementioned initial four-step procedure before applying MCFA. First, a conventional factor analysis with a single-factor model for the total sample covariance matrix is conducted to examine model specification and rough fit. Second, between-class variation is estimated using intraclass correlations (ICC) to ensure the amount of student well-being variation between classes. Note that ICCs range from 0 to 1, with the number close to 0 indicating less between-class variance. Third, a single factor structure using the pooled within-covariance matrix is estimated to explore individual-level variation. Fourth, a single factor structure using the between-class correlation matrix is estimated in order to explore the class-level structure (Muthén, 1994).

Initial analysis in four steps

Conventional factor analysis of $S_T$

All estimates are significant at the $p < .01$ level. Model fit indices are $\chi^2 (7) = 12.36$, $p < .01$, CFI = 1.000, TLI = 1.000, RMSEA = 0.017, and SRMR = 0.003 (Table 2). The result of the $S_T$ analysis shows a good fit. Additionally, all factor coefficients (loadings) were significantly different from zero ($p < .01$). Positive attitudes and emotions towards school had the highest loading ($b = 0.43$, followed by School worries ($b = 0.30$), Enjoyment in school ($b = 0.28$), Social problems in school ($b = 0.24$), Positive academic self-concept ($b = 0.22$), and Physical complaints ($b = 0.11$). Because the hierarchical nature of the data was ignored in these steps, using the fit indices and estimated parameters as part of the result is misleading. However, they do provide an indication of the model structure at the within level. The next step is to assess the between-class level information contained in the total matrix.

Estimation of between variations

The proportion between variations or ICCs of the six factors of student well-being ranged from 0.062 to 0.099, meaning that 6.2 to 9.9% of the variance of the variables results from between-group variation (Table 1). Physical complaints had the greatest between-class variability (ICC = 0.97), Positive Attitudes and Emotions toward school (ICC = 0.93), Enjoyment in school (ICC = 0.90), School worries (ICC = 0.62) and Positive academic self-concept (ICC = 0.049). These values indicated that there was sufficient between-class variability to ensure the multilevel analysis in the next step.

Estimation of within structure

In this step, the pool-within matrix $S_{pw}$ was analyzed. The proposed single-factor model is fitted to the pooled-within covariance matrix to determine whether the single-factor concept is appropriate for the pooled-within structure. All estimates are significant at the $p < .01$ level. Model fit indices are $\chi^2 (7) = 11.449$, $p < .01$, CFI = 1.000, TLI = 0.999, RMSEA = 0.015, and SRMR = 0.005 (Table 2). Using the $S_{pw}$ analysis in this step to explore the model structure provides a better fit and yields a superior approach than using the $S_T$ analysis from the first step does.
Estimation of between structures

The between structure is investigated in this step. The correlation matrix and an exploratory factor analysis are employed. A one-factor model shows that the eigenvalues range from 0.049 to 4.062 with the estimated loading ranging from 0.557 to 0.945. Thus, this model could be an acceptable fit for one-factor modeling of student well-being.  

Multilevel confirmatory factor analysis

Results of the two-level confirmatory factor analysis indicated a good fit of the model to the data. Those of the chi-square test were 17.473 with 12 degrees of freedom. The \( \chi^2 (df) \) was less than 2:1, which indicates a good fit. Moreover, according to other fit criteria, the RMSEA of 0.013 and CFI of 0.998 indicated an acceptable fit overall. The RMSEA fit indices at each level indicate that the fit at the student-level (within) part of the model was better than it was at the classroom-level (between) part (SRMR\(_{W} = 0.013\) vs. SRMR\(_{B} = 0.039\); see Table 2 for measures of fit).

At student level (within), all factor coefficients (loadings) were significantly different from zero (\( p < .01 \)). Positive attitudes and emotions towards school had the highest loading (\( b = 1.00 \)), followed by School worries (\( b = 0.67 \)), Enjoyment in school (\( b = 0.61 \)), Positive academic self-concept (\( b = 0.52 \)), Social problems in school (\( b = 0.51 \)), and Physical complaints (\( b = 0.20 \)). The within reliabilities were ranged from 0.02 to 0.73 (Table 3 and Figure 1).

At the classroom level (between), all factor coefficients (loadings) were significantly different from zero (\( p < .01 \)). Physical complaints (\( b = 1.00 \)) had the highest loading, followed by School worries (\( b = 0.91 \)), Social problems in school (\( b = 0.90 \)), Positive attitudes and emotions towards school (\( b = 0.57 \)), Enjoyment in school (\( b = 0.51 \)), and Positive academic self-concept (\( b = 0.43 \)). Table 3 displays the unstandardized factor loadings, standard errors, significant test, and reliabilities for both student and classroom levels. The between reliabilities ranged from 0.39 to 0.82 (Table 3 and Figure 1).

### Table 1. Descriptive statistics for student well-being (SWB) (n = 2,707)

<table>
<thead>
<tr>
<th>Factors</th>
<th>Mean</th>
<th>SD</th>
<th>Skewness</th>
<th>Kurtosis</th>
<th>ICC</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. Positive attitudes and emotions towards school</td>
<td>4.212</td>
<td>0.506</td>
<td>-0.736</td>
<td>0.832</td>
<td>0.093</td>
</tr>
<tr>
<td>2. Enjoyment in school</td>
<td>3.997</td>
<td>0.430</td>
<td>-0.403</td>
<td>0.272</td>
<td>0.900</td>
</tr>
<tr>
<td>3. Positive academic self-concept</td>
<td>3.589</td>
<td>0.465</td>
<td>0.023</td>
<td>0.295</td>
<td>0.049</td>
</tr>
<tr>
<td>4. (Absence of) Social problems in school</td>
<td>3.421</td>
<td>0.610</td>
<td>-0.337</td>
<td>0.298</td>
<td>0.097</td>
</tr>
<tr>
<td>5. (Absence of) School worries</td>
<td>3.729</td>
<td>0.753</td>
<td>-0.472</td>
<td>-0.059</td>
<td>0.062</td>
</tr>
<tr>
<td>6. (Absence of) Physical complaints</td>
<td>3.758</td>
<td>0.578</td>
<td>-0.575</td>
<td>0.689</td>
<td>0.099</td>
</tr>
</tbody>
</table>

Note: ICCs = Intraclass correlation coefficients

### Table 2. Summary of model fit for each model

<table>
<thead>
<tr>
<th>Model</th>
<th>( \chi^2 (df) )</th>
<th>P</th>
<th>CFI</th>
<th>TLI</th>
<th>RMSEA</th>
<th>SRMR</th>
</tr>
</thead>
<tbody>
<tr>
<td>Conventional CFA</td>
<td>12.36 (7)</td>
<td>0.0890</td>
<td>1.00</td>
<td>1.00</td>
<td>0.017</td>
<td>0.003</td>
</tr>
<tr>
<td>Pooled-within</td>
<td>11.449 (7)</td>
<td>0.1202</td>
<td>1.00</td>
<td>0.999</td>
<td>0.015</td>
<td>0.005</td>
</tr>
<tr>
<td>MCFA</td>
<td>17.473 (12)</td>
<td>0.1327</td>
<td>0.998</td>
<td>0.996</td>
<td>0.013</td>
<td>W: 0.013/B: 0.039</td>
</tr>
</tbody>
</table>

Note: CFI = norm comparative fit index; RMSEA = root mean square error of approximation; SRMR = standardized root mean square residual

### Table 3. Unstandardized coefficients for student well-being factors within- and between-class levels

<table>
<thead>
<tr>
<th>Factors</th>
<th>Within b</th>
<th>SE</th>
<th>T</th>
<th>R²</th>
<th>Between b</th>
<th>SE</th>
<th>t</th>
<th>R²</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. Positive attitudes and emotions towards school</td>
<td>1.00**</td>
<td>-</td>
<td>-</td>
<td>0.73</td>
<td>0.57</td>
<td>0.15</td>
<td>3.76**</td>
<td>0.39</td>
</tr>
<tr>
<td>2. Enjoyment in school</td>
<td>0.61</td>
<td>0.02</td>
<td>25.73**</td>
<td>0.38</td>
<td>0.51</td>
<td>0.12</td>
<td>4.38**</td>
<td>0.44</td>
</tr>
<tr>
<td>3. Positive academic self-concept</td>
<td>0.52</td>
<td>0.03</td>
<td>17.58**</td>
<td>0.22</td>
<td>0.43</td>
<td>0.11</td>
<td>3.91**</td>
<td>0.48</td>
</tr>
<tr>
<td>4. (Absence of) Social problems in school</td>
<td>0.51</td>
<td>0.04</td>
<td>13.84**</td>
<td>0.13</td>
<td>0.90</td>
<td>0.14</td>
<td>6.39**</td>
<td>0.61</td>
</tr>
<tr>
<td>5. (Absence of) School worries</td>
<td>0.67</td>
<td>0.04</td>
<td>18.01**</td>
<td>0.14</td>
<td>0.91</td>
<td>0.13</td>
<td>7.22**</td>
<td>0.65</td>
</tr>
<tr>
<td>6. (Absence of) Physical complaints</td>
<td>0.20</td>
<td>0.03</td>
<td>5.93**</td>
<td>0.02</td>
<td>1.00**</td>
<td>-</td>
<td>-</td>
<td>0.82</td>
</tr>
</tbody>
</table>

Note: ** p < .01, * program default
CONCLUSION

The application of MCFA provided two significant points of view with respect to our research questions. First, the factor structures for student well-being between student and class levels differ. While each individual student tended to mention his/her well-being differently in terms of his/her attitude and feelings, groups/classes of students tended to mention their physical health instead. Second, the reliabilities of student well-being factors varied differently between student and class levels, implying that the explained variance components of student well-being at each level were different as well. These two results were consistent and broaden our perspective on how research should examine student well-being at student and class levels.

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